## Statistical basics - A short overview

(discrete)

## The most important terms and definitions:

1) Expectation
2) Variance / Standard deviation
3) Sample variance
4) Covariance
5) Correlation coefficient
6) Independence vs. uncorrelation
7) Normal distribution and standard normal distribution

For comments and calculations see the appendix.

## To 1)

Generally: Let $g(X)$ be a unique function of the random variable $X$, then $g(X)$ is a random variable, too. In the discrete case we define the expectation of $g(X)$ as follows:

$$
\mathbf{E}[\mathbf{g}(\mathbf{X})]=\sum_{\mathrm{k}=1}^{\mathrm{n}} \mathbf{p}_{\mathrm{k}} \cdot \mathbf{g}\left(\mathbf{x}_{\mathrm{k}}\right) .
$$

The expectation of the random variable $X$ itself is obtained by setting $g(X)=X$ :

$$
\mu_{\mathbf{x}}=\mathbf{E}[\mathbf{X}]=\sum_{\mathbf{k}=1}^{\mathrm{n}} \mathbf{p}_{\mathrm{k}} \cdot \mathbf{x}_{\mathbf{k}} .
$$

Example, see appendix: ${ }^{i}$

In the lecture we considered the expectation of a portfolio with proportions $x_{i}$ invested in stock i. Further the return of stock $i$ is denoted by $r_{i}$.
The expected portfolio return is given by:

$$
\mathbf{E}[\mathbf{X}]=\mu=\sum_{\mathbf{i}=1}^{\mathbf{n}} \mathbf{x}_{\mathbf{i}} \cdot \mathbf{r}_{\mathbf{i}} .
$$

Notice: Therefore our portfolio return is already an expectation.

Expectation for $n$ variables (Assets), i.e. $\mathbf{E}[\mathbf{X}]=\mu=\mathbf{x}_{1} \cdot \mathbf{r}_{1}+\mathbf{x}_{2} \cdot \mathbf{r}_{2}+\ldots+\mathbf{x}_{\mathrm{n}} \cdot \mathbf{r}_{\mathbf{n}}$.
For instance
$\mathrm{n}=1: \mathbf{E}[\mathbf{a X}+\mathbf{b}]=\mathbf{a} \cdot \mathbf{E}[\mathbf{X}]+\mathbf{b}$, example, see appendix: ii
and $\mathrm{n}=2: \mathbf{E}[\mathbf{a X}+\mathbf{b Y}]=\mathbf{a} \cdot \mathbf{E}[\mathbf{X}]+\mathbf{b} \cdot \mathbf{E}[\mathbf{Y}]$, example, see appendix: iii

## To 2)

The variance is defined as the average quadratic deviation
$\mathbf{V}[\mathbf{X}]=\sigma^{\mathbf{2}}=\mathbf{E}[\mathbf{X}-\mathbf{E}(\mathbf{X})]^{2}$,
According to the "Theorem of Steiner" the variance can also be written as:
$\mathbf{V}[\mathbf{X}]=\sigma^{\mathbf{2}}=\mathbf{E}\left[\mathbf{X}^{2}\right]-[\mathbf{E}(\mathbf{X})]^{2}$.

The standard deviation is defined as the positive quadratic root of the variance:
$\sigma=\sqrt{\sigma^{2}}$.

## To 3)

Having the arithmetic mean of the distribution $\overline{\mathbf{x}}=\frac{1}{\mathbf{n}} \sum_{\mathrm{i}=1}^{\mathrm{n}} \mathbf{x}_{\mathbf{i}}$, the sample variance with $\mathrm{n}-1$ degrees of freedom is the following:

$$
\mathbf{S}^{2}=\hat{\sigma}^{2}=\frac{1}{\mathbf{n}-1} \sum_{\mathbf{i}=1}^{\mathbf{n}}\left(\mathbf{x}_{\mathrm{i}}-\overline{\mathbf{x}}\right)^{2}
$$

It holds $\mathbf{E}\left[\mathbf{S}^{2}\right]=\sigma^{2}$, therefore $\mathbf{S}^{2}$ is called an unbiased estimator of the variance $\sigma^{2}$.
Proof, see appendix: iv

## To 4)

The covariance measures the linear co-movement of $X$ and $Y$ :
$\operatorname{Cov}(\mathbf{X}, \mathbf{Y})=\sigma_{\mathbf{X}, \mathbf{Y}}=\mathbf{E}[(\mathbf{X}-\mathbf{E}(\mathbf{X})) \cdot(\mathbf{Y}-\mathbf{E}(\mathbf{Y}))]=\mathbf{E}[\mathbf{X Y}]-\mathbf{E}[\mathbf{X}] \cdot \mathbf{E}[\mathbf{Y}]$

## To 5)

The correlation coefficient is defined on $[-1,1]$ and has the following form:
$\operatorname{Corr}(\mathbf{X}, \mathbf{Y})=\rho_{\mathbf{X}, \mathbf{Y}}=\frac{\operatorname{Cov}(\mathbf{X}, \mathbf{Y})}{\sqrt{\mathbf{V}(\mathbf{X}) \cdot \mathbf{V}(\mathbf{Y})}}=\frac{\sigma_{\mathbf{X}, \mathbf{Y}}}{\sigma_{\mathbf{X}} \cdot \sigma_{\mathbf{Y}}}$.

With 5) we receive this equation (used in the tutorial) :
$\boldsymbol{\operatorname { C o v }}(\mathbf{X}, \mathbf{Y})=\boldsymbol{\sigma}_{\mathbf{X}, \mathbf{Y}}=\boldsymbol{\rho}_{\mathbf{X}, \mathbf{Y}} \cdot \sqrt{\mathbf{V}(\mathbf{X}) \cdot \mathbf{V}(\mathbf{Y})} \Leftrightarrow \mathbf{C o v}(\mathbf{X}, \mathbf{Y})=\rho_{\mathbf{X}, \mathbf{Y}} \cdot \sigma_{\mathbf{X}} \cdot \sigma_{\mathbf{Y}}$

## To 6)

Generally, you cannot take the following implication: an uncorrelated random variable is also independent. This holds only for symmetric distributions, like the normal distribution.

## To 7)

A random variable X is called normal distributed, for short: $\mathrm{X} \sim \mathrm{N}\left(\mu, \sigma^{2}\right)$, if it has a normal density function with parameters $\mu$ and $\sigma^{2}$.

Standard Normal density


A normal distribution with $\mu=0$ and $\sigma^{2}=1$ is called a standard normal distribution $\mathrm{N}(0,1)$.
The distribution has the density function $\mathbf{n}(\mathbf{x})=\frac{1}{\sqrt{2 \pi}} \mathbf{e}^{-\frac{\mathbf{x}^{2}}{2}}$, as plotted in the graph above:

Some properties of the normal distribution:

- unimodal distribution
- symmetric distribution with maximum at $x=\mu$
- points of inflexion at $\mathrm{x}=\mu \pm \sigma$
- $\mathrm{E}(\mathrm{X})=\mu, \operatorname{Var}(\mathrm{X})=\sigma^{2}$

You can transform any normal distribution into a standard normal distribution.
For $\mathrm{X} \sim \mathrm{N}\left(\mu, \sigma^{2}\right)$-distributed random variable, $\mathbf{U}:=\frac{\mathbf{x}-\mu}{\sigma}$ is a standard normal distributed random variable.

For that reason all calculations (probabilities, quantiles, a.s.o.) can be done based on a standard normal distribution and you do not need to do calculation for each $\mathrm{N}\left(\mu, \sigma^{\mathbf{2}}\right)$-distributed random variable.

Let $\mathbf{x}_{\mathrm{p}}$ be the quantile of order p of a $\mathrm{N}\left(\mu, \sigma^{2}\right)$-distribution and $\lambda_{\mathrm{p}}$ the quantile of order p of a $\mathrm{N}(0,1)$ distribution. Then:

$$
\mathbf{x}_{\mathrm{p}}=\mu+\lambda_{\mathrm{p}} \cdot \sigma, \forall \mathbf{p} \in(0,1) .
$$

By considering a normal distribution the central coverage interval lies symmetrically around the mean. Having these $N(0,1)$ quantiles you can determine the coverage interval for a probability $1-\alpha$ for a $N($ $\mu, \sigma^{2}$ )-distributed random variable X as

$$
\mathbf{P}\left(\mu-\lambda_{1-\frac{\alpha}{2}} \cdot \sigma \leq \mathbf{X} \leq \mu+\lambda_{1-\frac{\alpha}{2}} \cdot \sigma\right)=1-\alpha, \text { with } \lambda_{\mathbf{p}}=-\lambda_{1-\mathbf{p}}
$$

The other way round you can calculate out of the quantiles, the corresponding probability of the realization lying between the $\mu \pm \lambda_{1-\frac{\alpha}{2}} \sigma$ interval. That probability can be interpreted as the relative frequency.

For $\lambda_{1-\frac{\alpha}{2}}=\mathrm{K}$ we receive the relative frequency for the $\mathrm{K}^{\star} \sigma$-bands. For instance:
K=1: $\quad P(\mu-\sigma \leq X \leq \mu+\sigma)=0.6827$
i.e. approximately $68 \%$ of all normal realisations lie within the band $\mu \pm \sigma$.
$\mathrm{K}=2$ :
$\mathbf{P}(\mu-2 \sigma \leq \mathbf{X} \leq \mu+2 \sigma)=0,9545$
i.e. approximately $95 \%$ of all normal realisations lie within the band $\mu \pm 2 \sigma$.

$$
\mathrm{K}=3: \quad \mathbf{P}(\mu-3 \sigma \leq \mathbf{X} \leq \mu+3 \sigma)=0,9973
$$

i.e. approximately $99.7 \%$ of all normal realisations lie within the band $\mu \pm 3 \sigma$.

## Appendix:

i E.g. expectation of a discrete distribution function $\mathrm{P}\left(\mathrm{X}=\mathrm{x}_{\mathrm{i}}\right)$ :

$$
\mathbf{E}[\mathbf{X}]=\sum_{\mathrm{i}=1}^{\mathrm{n}} \mathbf{x}_{\mathrm{i}} \cdot \mathbf{p}\left(\mathbf{x}_{\mathrm{i}}\right)=\sum_{\mathrm{i}=1}^{\mathrm{n}} \mathbf{x}_{\mathrm{i}} \cdot \mathbf{P}\left(\mathbf{X}=\mathbf{x}_{\mathrm{i}}\right)
$$

example (chapter risk and return, slide 8):
Assumption: A discrete random variable X with density function $\mathrm{f}(\mathrm{x})$, and constants a and b .
Proposition: $\mathbf{E}[\mathbf{a X}+\mathbf{b}]=\mathbf{a} \cdot \mathbf{E}[\mathbf{X}]+\mathbf{b}$
Prove: $\quad \mathbf{E}[\mathbf{a X}+\mathbf{b}]=\sum_{\mathrm{i}=1}^{\mathrm{m}}\left(\mathbf{a} \mathbf{x}_{\mathrm{i}}+\mathbf{b}\right) \cdot \mathbf{f}\left(\mathbf{x}_{\mathrm{i}}\right)$

$$
\begin{aligned}
& =\sum_{i} \mathbf{a} \mathbf{x}_{\mathbf{i}} \mathbf{f}\left(\mathbf{x}_{\mathrm{i}}\right)+\mathbf{b f}\left(\mathbf{x}_{\mathbf{i}}\right) \\
& =\mathbf{a} \sum_{\mathbf{i}} \mathbf{x}_{\mathbf{i}} \mathbf{f}\left(\mathbf{x}_{\mathbf{i}}\right)+\mathbf{b} \sum_{\mathbf{i}} \mathbf{f}\left(\mathbf{x}_{\mathbf{i}}\right) \\
& =\mathbf{a E}[\mathbf{X}]+\mathbf{b} .
\end{aligned}
$$

q.e.d.

You can show $\mathbf{V}[\mathbf{a X}+\mathbf{b}]=\mathbf{a}^{\mathbf{2}} \cdot \mathbf{V}[\mathbf{X}]$ in the same way.

As well as for the variance: $\mathbf{V}[\mathbf{a X}+\mathbf{b Y}]=\mathbf{a}^{\mathbf{2}} \cdot \mathbf{V}[\mathbf{X}]+\mathbf{b}^{\mathbf{2}} \mathbf{V}[\mathbf{Y}]+2 \cdot \mathbf{a} \cdot \mathbf{b} \cdot \mathbf{C o v}(\mathbf{X}, \mathbf{Y})$ $\Leftrightarrow \mathbf{V}[\mathbf{a X}+\mathbf{b Y}]=\mathbf{a}^{2} \cdot \mathbf{V}[\mathbf{X}]+\mathbf{b}^{2} \mathbf{V}[\mathbf{Y}]+2 \cdot \mathbf{a} \cdot \mathbf{b} \cdot \sigma_{\mathbf{X}} \cdot \sigma_{\mathbf{Y}} \cdot \rho_{\mathbf{X} ; \mathbf{Y}}$
iv
calculation :
$\mathbf{E}\left[\mathbf{S}^{\mathbf{2}}\right]=\mathbf{E}\left[\frac{1}{\mathbf{n}-1} \sum_{\mathrm{i}=1}^{\mathrm{n}}\left(\mathbf{x}_{\mathbf{i}}-\overline{\mathbf{x}}\right)^{2}\right]$
$=\frac{\mathbf{n}}{\mathbf{n}-1} \cdot \frac{1}{\mathbf{n}} \mathbf{E}\left[\sum_{\mathbf{i}=1}^{\mathbf{n}}\left(\mathbf{x}_{\mathbf{i}}-\overline{\mathbf{x}}\right)^{2}\right]$
$=\frac{\mathbf{n}}{\mathbf{n}-1} \cdot \frac{1}{\mathbf{n}} \sum_{\mathbf{i}} \mathbf{E}\left[\left(\mathbf{x}_{\mathbf{i}}-\overline{\mathbf{x}}\right)^{2}\right]=\frac{\mathbf{n}}{\mathbf{n}-1} \cdot \frac{1}{\mathbf{n}} \sum_{\mathbf{i}} \mathbf{E}\left[\left(\left\{\mathbf{x}_{\mathbf{i}}-\mu\right\}-\{\overline{\mathbf{x}}-\mu\}\right)^{2}\right]$
with the binomial formula and $\mathbf{E}[(\overline{\mathbf{x}}-\mu)]=\frac{\sigma}{\mathbf{n}}$ it follows:

$$
\begin{aligned}
& \mathbf{E}\left[\mathbf{S}^{2}\right]=\frac{\mathbf{n}}{\mathbf{n}-1} \cdot \frac{1}{\mathbf{n}}\left[\mathbf{n} \cdot \sigma^{2}+\frac{\mathbf{n} \sigma^{2}}{\mathbf{n}}-2 \cdot \mathbf{E}\left[(\overline{\mathbf{x}}-\mu) \sum\left(\mathbf{x}_{\mathbf{i}}-\mu\right)\right]\right] \\
& =\frac{\mathbf{n}}{\mathbf{n}-1}\left[\sigma^{2}+\frac{\sigma^{2}}{\mathbf{n}}-\frac{2 \sigma^{2}}{\mathbf{n}}\right]=\frac{\mathbf{n}}{\mathbf{n}-1}\left[1+\frac{1}{\mathbf{n}}-\frac{2}{\mathbf{n}}\right] \sigma^{2}=\sigma^{2} .
\end{aligned}
$$

q.e.d.

