## Statistical basics - A short overview

# (discrete)

The most important terms and definitions:

- 1) Expectation
- 2) Variance / Standard deviation
- 3) Sample variance
- 4) Covariance
- 5) Correlation coefficient
- 6) Independence vs. uncorrelation
- 7) Normal distribution and standard normal distribution

For comments and calculations see the appendix.

## <u>To 1)</u>

Generally: Let g(X) be a unique function of the random variable X, then g(X) is a random variable, too. In the discrete case we define the expectation of g(X) as follows:

$$\mathbf{E}[\mathbf{g}(\mathbf{X})] = \sum_{k=1}^{n} \mathbf{p}_{k} \cdot \mathbf{g}(\mathbf{x}_{k})$$

The expectation of the random variable X itself is obtained by setting g(X)=X:

$$\mu_{\mathbf{X}} = \mathbf{E}[\mathbf{X}] = \sum_{k=1}^{n} \mathbf{p}_{k} \cdot \mathbf{x}_{k}$$

Example, see appendix: i

In the lecture we considered the expectation of a portfolio with proportions  $x_i$  invested in stock i. Further the return of stock i is denoted by  $r_i$ .

The expected portfolio return is given by:

$$\mathbf{E}[\mathbf{X}] = \mu = \sum_{i=1}^{n} \mathbf{x}_{i} \cdot \mathbf{r}_{i} .$$

Notice: Therefore our portfolio return is already an expectation.

Expectation for n variables (Assets), i.e.  $\mathbf{E}[\mathbf{X}] = \mu = \mathbf{x}_1 \cdot \mathbf{r}_1 + \mathbf{x}_2 \cdot \mathbf{r}_2 + \dots + \mathbf{x}_n \cdot \mathbf{r}_n$ .

For instance

n=1:  $\mathbf{E}[\mathbf{aX} + \mathbf{b}] = \mathbf{a} \cdot \mathbf{E}[\mathbf{X}] + \mathbf{b}$ , example, see appendix: <sup>ii</sup>

and n=2:  $\mathbf{E}[\mathbf{aX} + \mathbf{bY}] = \mathbf{a} \cdot \mathbf{E}[\mathbf{X}] + \mathbf{b} \cdot \mathbf{E}[\mathbf{Y}]$ , example, see appendix: <sup>iii</sup>

#### <u>To 2)</u>

The variance is defined as the average quadratic deviation

$$\mathbf{V}[\mathbf{X}] = \sigma^2 = \mathbf{E}[\mathbf{X} - \mathbf{E}(\mathbf{X})]^2 ,$$

According to the "Theorem of Steiner" the variance can also be written as:

$$\mathbf{V}[\mathbf{X}] = \sigma^2 = \mathbf{E}[\mathbf{X}^2] - [\mathbf{E}(\mathbf{X})]^2 \ .$$

The standard deviation is defined as the positive quadratic root of the variance:

$$\sigma = \sqrt{\sigma^2}$$
.

#### <u>To 3)</u>

Having the arithmetic mean of the distribution  $\overline{\mathbf{x}} = \frac{1}{n} \sum_{i=1}^{n} \mathbf{x}_{i}$ , the sample variance with n-1 degrees of

freedom is the following:

$$\mathbf{S}^2 = \hat{\sigma}^2 = \frac{1}{\mathbf{n} - 1} \sum_{i=1}^{\mathbf{n}} (\mathbf{x}_i - \overline{\mathbf{x}})^2$$

It holds  $\mathbf{E}[\mathbf{S}^2] = \sigma^2$ , therefore S<sup>2</sup> is called an unbiased estimator of the variance  $\sigma^2$ . Proof, see appendix: <sup>iv</sup>

### <u>To 4)</u>

The covariance measures the linear co-movement of X and Y:

 $Cov(X, Y) = \sigma_{X,Y} = E[(X - E(X)) \cdot (Y - E(Y))] = E[XY] - E[X] \cdot E[Y]$ 

### <u>To 5)</u>

The correlation coefficient is defined on [-1,1] and has the following form:

$$\operatorname{Corr}(\mathbf{X}, \mathbf{Y}) = \rho_{\mathbf{X}, \mathbf{Y}} = \frac{\operatorname{Cov}(\mathbf{X}, \mathbf{Y})}{\sqrt{\operatorname{V}(\mathbf{X}) \cdot \operatorname{V}(\mathbf{Y})}} = \frac{\sigma_{\mathbf{X}, \mathbf{Y}}}{\sigma_{\mathbf{X}} \cdot \sigma_{\mathbf{Y}}}$$

With 5) we receive this equation (used in the tutorial) :

$$\operatorname{Cov}(\mathbf{X},\mathbf{Y}) = \boldsymbol{\sigma}_{\mathbf{X},\mathbf{Y}} = \boldsymbol{\rho}_{\mathbf{X},\mathbf{Y}} \cdot \sqrt{\mathbf{V}(\mathbf{X}) \cdot \mathbf{V}(\mathbf{Y})} \Leftrightarrow \operatorname{Cov}(\mathbf{X},\mathbf{Y}) = \boldsymbol{\rho}_{\mathbf{X},\mathbf{Y}} \cdot \boldsymbol{\sigma}_{\mathbf{X}} \cdot \boldsymbol{\sigma}_{\mathbf{Y}}$$

### <u>To 6)</u>

Generally, you cannot take the following implication: an uncorrelated random variable is also independent. This holds only for symmetric distributions, like the normal distribution.

#### <u>To 7)</u>

A random variable X is called normal distributed, for short: X ~ N ( $\mu$  ,  $\sigma^2$ ), if it has a normal density function with parameters  $\mu$  and  $\sigma^2$ .



#### Standard Normal density

A normal distribution with  $\mu = 0$  and  $\sigma^2 = 1$  is called a standard normal distribution N(0,1).

The distribution has the density function  $\mathbf{n}(\mathbf{x}) = \frac{1}{\sqrt{2\pi}} \mathbf{e}^{-\frac{\mathbf{x}^2}{2}}$ , as plotted in the graph above:

Some properties of the normal distribution:

- unimodal distribution
- $\cdot$  symmetric distribution with maximum at x =  $\mu$
- · points of inflexion at x =  $\mu \pm \sigma$

$$\cdot$$
 E (X) =  $\mu$ , Var (X) =  $\sigma^2$ 

You can transform any normal distribution into a standard normal distribution.

For X ~ N ( $\mu$ ,  $\sigma^2$ )-distributed random variable,  $\mathbf{U} := \frac{\mathbf{x} - \mu}{\sigma}$  is a standard normal distributed random

variable.

For that reason all calculations (probabilities, quantiles, a.s.o.) can be done based on a standard normal distribution and you do not need to do calculation for each N ( $\mu$ ,  $\sigma^2$ )-distributed random variable.

Let  $\mathbf{x}_{p}$  be the quantile of order p of a N ( $\mu$ ,  $\sigma^{2}$ )-distribution and  $\lambda_{p}$  the quantile of order p of a N(0, 1)distribution. Then:

$$\mathbf{x}_{\mathbf{p}} = \mu + \lambda_{\mathbf{p}} \cdot \boldsymbol{\sigma}, \forall \mathbf{p} \in (0,1).$$

By considering a normal distribution the central coverage interval lies symmetrically around the mean. Having these N(0,1) quantiles you can determine the coverage interval for a probability  $1-\alpha$  for a N( $\mu$ ,  $\sigma^2$ )-distributed random variable X as

$$\mathbf{P}(\mu - \lambda_{1-\frac{\alpha}{2}} \cdot \sigma \leq \mathbf{X} \leq \mu + \lambda_{1-\frac{\alpha}{2}} \cdot \sigma) = 1 - \alpha \text{ , with } \lambda_{\mathbf{p}} = -\lambda_{1-\mathbf{p}}$$

The other way round you can calculate out of the quantiles, the corresponding probability of the realization lying between the  $\mu \pm \lambda_{1-\frac{\alpha}{2}} \sigma$  interval. That probability can be interpreted as the relative frequency.

frequency.

For  $\lambda_{1-\frac{\alpha}{2}} = K$  we receive the relative frequency for the K<sup>\*</sup> $\sigma$ -bands. For instance: K=1:  $P(\mu - \sigma \le X \le \mu + \sigma) = 0.6827$ 

i.e. approximately 68% of all normal realisations lie within the band  $\mu \pm \sigma$ .

K=2: 
$$P(\mu - 2\sigma \le X \le \mu + 2\sigma) = 0,9545$$

i.e. approximately 95% of all normal realisations lie within the band  $\,\mu^{}$   $\pm 2\,\sigma$  .

K=3: 
$$P(\mu - 3\sigma \le X \le \mu + 3\sigma) = 0,9973$$

i.e. approximately 99.7% of all normal realisations lie within the band  $\,\mu\,\,\pm 3\,\sigma$  .

### Appendix:

i

ii

*E.g.* expectation of a discrete distribution function  $P(X = x_i)$ :

$$\mathbf{E}[\mathbf{X}] = \sum_{i=1}^{n} \mathbf{x}_{i} \cdot \mathbf{p}(\mathbf{x}_{i}) = \sum_{i=1}^{n} \mathbf{x}_{i} \cdot \mathbf{P}(\mathbf{X} = \mathbf{x}_{i})$$

*example* (chapter risk and return, slide 8):

Assumption: A discrete random variable X with density function f(x), and constants a and b.

Proposition: 
$$\mathbf{E}[\mathbf{aX} + \mathbf{b}] = \mathbf{a} \cdot \mathbf{E}[\mathbf{X}] + \mathbf{b}$$
  
Prove:  $\mathbf{E}[\mathbf{aX} + \mathbf{b}] = \sum_{i=1}^{m} (\mathbf{ax}_i + \mathbf{b}) \cdot \mathbf{f}(\mathbf{x}_i)$   
 $= \sum_i \mathbf{ax}_i \mathbf{f}(\mathbf{x}_i) + \mathbf{bf}(\mathbf{x}_i)$   
 $= \mathbf{a}\sum_i \mathbf{x}_i \mathbf{f}(\mathbf{x}_i) + \mathbf{b}\sum_i \mathbf{f}(\mathbf{x}_i)$   
 $= \mathbf{a}\mathbf{E}[\mathbf{X}] + \mathbf{b}.$ 

q.e.d.

calculation :

You can show  $V[aX + b] = a^2 \cdot V[X]$  in the same way.

As well as for the variance:  $V[aX + bY] = a^2 \cdot V[X] + b^2 V[Y] + 2 \cdot a \cdot b \cdot Cov(X, Y)$  $\Leftrightarrow V[aX + bY] = a^2 \cdot V[X] + b^2 V[Y] + 2 \cdot a \cdot b \cdot \sigma_X \cdot \sigma_Y \cdot \rho_{X;Y}$ 

iii

iv

$$\begin{split} \mathbf{E}[\mathbf{S}^2] &= \mathbf{E}\left[\frac{1}{\mathbf{n}-1}\sum_{i=1}^{\mathbf{n}}\left(\mathbf{x}_i - \overline{\mathbf{x}}\right)^2\right] \\ &= \frac{\mathbf{n}}{\mathbf{n}-1} \cdot \frac{1}{\mathbf{n}} \mathbf{E}\left[\sum_{i=1}^{\mathbf{n}}\left(\mathbf{x}_i - \overline{\mathbf{x}}\right)^2\right] \\ &= \frac{\mathbf{n}}{\mathbf{n}-1} \cdot \frac{1}{\mathbf{n}} \sum_{i} \mathbf{E}[(\mathbf{x}_i - \overline{\mathbf{x}})^2] = \frac{\mathbf{n}}{\mathbf{n}-1} \cdot \frac{1}{\mathbf{n}} \sum_{i} \mathbf{E}[(\{\mathbf{x}_i - \mu\} - \{\overline{\mathbf{x}} - \mu\})^2] \end{split}$$

with the binomial formula and  $\mathbf{E}[(\bar{\mathbf{x}} - \mu)] = \frac{\sigma}{n}$  it follows:

$$\mathbf{E}[\mathbf{S}^{2}] = \frac{\mathbf{n}}{\mathbf{n}-1} \cdot \frac{1}{\mathbf{n}} \left[ \mathbf{n} \cdot \sigma^{2} + \frac{\mathbf{n}\sigma^{2}}{\mathbf{n}} - 2 \cdot \mathbf{E}[(\overline{\mathbf{x}} - \mu)\sum(\mathbf{x}_{i} - \mu)] \right]$$
$$= \frac{\mathbf{n}}{\mathbf{n}-1} \left[ \sigma^{2} + \frac{\sigma^{2}}{\mathbf{n}} - \frac{2\sigma^{2}}{\mathbf{n}} \right] = \frac{\mathbf{n}}{\mathbf{n}-1} \left[ 1 + \frac{1}{\mathbf{n}} - \frac{2}{\mathbf{n}} \right] \sigma^{2} = \sigma^{2}.$$

q.e.d.